

Much of this is taken almost directly from both the Stata manual on `ivregress` and lecture notes from Sally Stearn (UNC). We are interested in estimating the effect of housing value on rent, however as the Stata manual states we are worried about random shocks (our omitted variable) affecting both housing values and rent. To counteract this, we use average family income and census region as instruments for housing value. The data is state level averages from 1980.

```
. use https://www.stata-press.com/data/r18/hsng
(1980 Census housing data)
```

```
. sum
```

Variable	Obs	Mean	Std. Dev.	Min	Max
state	0				
division	50	5.12	2.560612	1	9
region	50	2.66	1.061574	1	4
pop	50	4518149	4715038	401851	2.37e+07
popgrow	50	16.29	14.38729	-3.6	63.8
popden	50	1543.72	2213.894	7	9862
pcturban	50	66.94913	14.40956	33.77319	91.29498
faminc	50	19499.92	2617.218	14591	28395
hsng	50	1762686	1847308	162825	9279036
hsnggrow	50	35.65202	19.30491	9.015595	97.00565
hsngval	50	48484	15770.24	31100	119400
rent	50	234.76	35.35335	180	368
reg1	50	.18	.3880879	0	1
reg2	50	.24	.4314191	0	1
reg3	50	.32	.4712121	0	1
reg4	50	.26	.4430875	0	1

Start with the OLS regression (we are worried that this might be biased due to random shocks).

```
. reg rent hsngval pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	40983.5269	2	20491.7635	F(2, 47)	=	47.54
Residual	20259.5931	47	431.055172	Prob > F	=	0.0000
Total	61243.12	49	1249.85959	R-squared	=	0.6692
				Adj R-squared	=	0.6551
				Root MSE	=	20.762

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hsngval	.0015205	.0002276	6.68	0.000	.0010627 .0019784
pcturban	.5248216	.2490782	2.11	0.040	.0237408 1.025902
_cons	125.9033	14.18537	8.88	0.000	97.36603 154.4406

Store the estimates to be used later in the Durban-Wu-Hausman test.

```
. est store my_OLS
```

Before we run our IV, we will test the power of the instrument by manually running the first stage. Note that we include the instruments as well as any control variables that will also be included in the second stage.

```
. reg hsngval faminc reg2-reg4 pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	8.4187e+09	5	1.6837e+09	F(5, 44)	=	19.66
Residual	3.7676e+09	44	85626930.6	Prob > F	=	0.0000
Total	1.2186e+10	49	248700555	R-squared	=	0.6908
				Adj R-squared	=	0.6557
				Root MSE	=	9253.5

hsngval	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
---------	-------	-----------	---	------	----------------------

-----+-----						
faminc		2.731324	.6818931	4.01	0.000	1.357058 4.105589
reg2		-5095.038	4122.112	-1.24	0.223	-13402.61 3212.533
reg3		-1778.05	4072.691	-0.44	0.665	-9986.019 6429.919
reg4		13413.79	4048.141	3.31	0.002	5255.296 21572.28
pcturban		182.2201	115.0167	1.58	0.120	-49.58092 414.0211
_cons		-18671.87	11995.48	-1.56	0.127	-42847.17 5503.439

The F-test can be run after the first stage has been estimated. The F-statistic is above 10, which is thought of by many as the minimum threshold to not worry about a weak instrument problem.

```
. testparm faminc reg2-reg4

( 1)  faminc = 0
( 2)  reg2 = 0
( 3)  reg3 = 0
( 4)  reg4 = 0

      F( 4, 44) = 13.30
      Prob > F = 0.0000
```

Next we estimate the IV. Note that we are using two stage least squares (2sls). The outcome in the second stage is included first, then any control variables that will be included in both the second and first stage, with the first stage in the parenthesis. The endogenous variable is included to the left of the = and the instruments that will be omitted from the second stage are included to the right. I have also added the "first" option that reports the results of the first stage in the output.

```
. ivregress 2sls rent pcturban (hsngval = faminc reg2-reg4), first
```

First-stage regressions

				Number of obs	=	50
				F(5, 44)	=	19.66
				Prob > F	=	0.0000
				R-squared	=	0.6908
				Adj R-squared	=	0.6557
				Root MSE	=	9253.4821

hsngval		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
pcturban		182.2201	115.0167	1.58	0.120	-49.58092 414.0211
faminc		2.731324	.6818931	4.01	0.000	1.357058 4.105589
reg2		-5095.038	4122.112	-1.24	0.223	-13402.61 3212.533
reg3		-1778.05	4072.691	-0.44	0.665	-9986.019 6429.919
reg4		13413.79	4048.141	3.31	0.002	5255.296 21572.28
_cons		-18671.87	11995.48	-1.56	0.127	-42847.17 5503.439

Instrumental variables (2SLS) regression

```
Number of obs = 50
Wald chi2(2) = 90.76
Prob > chi2 = 0.0000
R-squared = 0.5989
Root MSE = 22.166
```

rent		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
hsngval		.0022398	.0003284	6.82	0.000	.0015961 .0028836
pcturban		.081516	.2987652	0.27	0.785	-.504053 .667085
_cons		120.7065	15.22839	7.93	0.000	90.85942 150.5536

Instrumented: hsngval

Instruments: pcturban faminc reg2 reg3 reg4

Store the estimates to use later in the Durbin-Wu-Hausman test.

```
. est store my_IV
```

After running the IV you can use this command to estimate several first stage weak instrument tests. Note that the F is the same as the F-test we conducted above. We also have the partial R² which is another common test for the strength of an instrument. In this case 0.547 is quite large. Rather than using the critical value of 10 for the F-test, the output includes some different critical values depending on the amount of relative bias that the researcher is willing to accept.

```
. estat firststage
```

First-stage regression summary statistics

Variable	R-sq.	Adjusted R-sq.	Partial R-sq.	F(4,44)	Prob > F
hsngval	0.6908	0.6557	0.5473	13.2978	0.0000

Minimum eigenvalue statistic = 13.2978

Critical Values # of endogenous regressors: 1
Ho: Instruments are weak # of excluded instruments: 4

	5%	10%	20%	30%
2SLS relative bias	16.85	10.27	6.71	5.34
2SLS Size of nominal 5% Wald test	24.58	13.96	10.26	8.31
LIML Size of nominal 5% Wald test	5.44	3.87	3.30	2.98

The Durban-Wu-Hausman tests if the coefficients from a consistent estimator (IV) are different from those estimated from a more efficient but potentially biased estimator (OLS). The Chi² indicates that the coefficients are different, meaning X is endogenous and we should be using IV. However, there is a common problem with this test where the V_b-V_B is not positive definite. I want to reiterate that to conduct this test and the variants of the test, we need to assume that the instrument is strong and not correlated with the error term - so this is NOT a test of an instrument's suitability.

```
. hausman my_IV my_OLS, sigmamore
```

Note: the rank of the differenced variance matrix (1) does not equal the number of coefficients being tested (2); be sure this is what you expect, or there may be problems computing the test. Examine the output of your estimators for anything unexpected and possibly consider scaling your variables so that the coefficients are on a similar scale.

	(b)	(B)	(b-B)	sqrt(diag(V _b -V _B))
	my_IV	my_OLS	Difference	S.E.
hsngval	.0022398	.0015205	.0007193	.000207
pcturban	.081516	.5248216	-.4433056	.1275655

b = consistent under Ho and Ha; obtained from ivregress
B = inconsistent under Ha, efficient under Ho; obtained from regress

Test: Ho: difference in coefficients not systematic

```
chi2(1) = (b-B)'[(Vb-VB)-1](b-B)
        = 12.08
Prob>chi2 = 0.0005
(Vb-VB is not positive definite)
```

There are a few variants of the Hausman test. The first is to predict the endogenous variable using the first stage regression and include in the second stage regression. The second is to include the residual instead. I calculate both the predicted X and residual below.

```
. reg hsnngval faminc reg2-reg4 pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	8.4187e+09	5	1.6837e+09	F(5, 44)	=	19.66
Residual	3.7676e+09	44	85626930.6	Prob > F	=	0.0000
				R-squared	=	0.6908
				Adj R-squared	=	0.6557
Total	1.2186e+10	49	248700555	Root MSE	=	9253.5

hsnngval	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	2.731324	.6818931	4.01	0.000	1.357058	4.105589
reg2	-5095.038	4122.112	-1.24	0.223	-13402.61	3212.533
reg3	-1778.05	4072.691	-0.44	0.665	-9986.019	6429.919
reg4	13413.79	4048.141	3.31	0.002	5255.296	21572.28
pcturban	182.2201	115.0167	1.58	0.120	-49.58092	414.0211
_cons	-18671.87	11995.48	-1.56	0.127	-42847.17	5503.439

Predicted X

```
. predict hsnng_hat
(option xb assumed; fitted values)
```

Residual

```
. predict FSres, res

. su hsnngval hsnng_hat FSres
```

Variable	Obs	Mean	Std. Dev.	Min	Max
hsnngval	50	48484	15770.24	31100	119400
hsnng_hat	50	48484	13107.67	28025.76	84022.63
FSres	50	.0000372	8768.665	-10503.84	46755.89

Before conducting those tests, I want to show that we get the same coefficient as IV when we estimate the second stage using the predicted X from the first stage, however our standard errors are not correct. Compare this...

```
. reg rent hsnng_hat pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	44592.8127	2	22296.4063	F(2, 47)	=	62.94
Residual	16650.3073	47	354.261858	Prob > F	=	0.0000
				R-squared	=	0.7281
				Adj R-squared	=	0.7166
Total	61243.12	49	1249.85959	Root MSE	=	18.822

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsnng_hat	.0022398	.0002789	8.03	0.000	.0016788	.0028009
pcturban	.0815159	.2536953	0.32	0.749	-.4288532	.5918851
_cons	120.7065	12.93113	9.33	0.000	94.69244	146.7206

To this.

```
. ivregress 2sls rent pcturban (hsnngval = faminc reg2-reg4)
```

Instrumental variables (2SLS) regression	Number of obs	=	50
	Wald chi2(2)	=	90.76
	Prob > chi2	=	0.0000
	R-squared	=	0.5989
	Root MSE	=	22.166

rent	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
hsnngval	.0022398	.0003284	6.82	0.000	.0015961	.0028836
pcturban	.081516	.2987652	0.27	0.785	-.504053	.667085
_cons	120.7065	15.22839	7.93	0.000	90.85942	150.5536

```
-----
Instrumented:  hsnngval
Instruments:   pcturban faminc reg2 reg3 reg4
```

Now back to our first test, estimate the second stage including the X and predicted X. The null is that the predicted X should not provide different variation from X (or that the coefficient is 0) and the alternative is that it does. We should only get a different value if the $cov(X, e)$ does not equal 0 (or X is endogenous). In this case we reject the null that there is no different coefficient (see the t-stat on the hsnng_hat variable).

```
. reg rent hsnngval hsnng_hat pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	46189.152	3	15396.384	F(3, 46)	=	47.05
Residual	15053.968	46	327.260173	Prob > F	=	0.0000
				R-squared	=	0.7542
				Adj R-squared	=	0.7382
Total	61243.12	49	1249.85959	Root MSE	=	18.09

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsnngval	.0006509	.0002947	2.21	0.032	.0000577	.0012442
hsnng_hat	.0015889	.0003984	3.99	0.000	.000787	.0023908
pcturban	.0815159	.2438355	0.33	0.740	-.4092994	.5723313
_cons	120.7065	12.42856	9.71	0.000	95.68912	145.7239

You come to the same conclusion using the residual from the first stage.

```
. reg rent hsnngval FSres pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	46189.1513	3	15396.3838	F(3, 46)	=	47.05
Residual	15053.9687	46	327.26019	Prob > F	=	0.0000
				R-squared	=	0.7542
				Adj R-squared	=	0.7382
Total	61243.12	49	1249.85959	Root MSE	=	18.09

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsnngval	.0022398	.0002681	8.36	0.000	.0017003	.0027794
FSres	-.0015889	.0003984	-3.99	0.000	-.0023908	-.000787
pcturban	.081516	.2438355	0.33	0.740	-.4092994	.5723313
_cons	120.7065	12.42856	9.71	0.000	95.68912	145.7239

The last variant is to include the instruments in the second stage. The null here is that the instruments should have a 0 coefficient. We find that is not the case. Note that I used an F-test since we have multiple instruments.

```
. reg rent hsnngval faminc reg2-reg4 pcturban
```

Source	SS	df	MS	Number of obs	=	50
Model	51734.9429	6	8622.49049	F(6, 43)	=	38.99
Residual	9508.17708	43	221.120397	Prob > F	=	0.0000
				R-squared	=	0.8447
				Adj R-squared	=	0.8231
Total	61243.12	49	1249.85959	Root MSE	=	14.87

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsnngval	.0006509	.0002423	2.69	0.010	.0001624	.0011395
faminc	.0087876	.0012801	6.86	0.000	.0062061	.0113691
reg2	-14.86383	6.738152	-2.21	0.033	-28.45261	-1.275052
reg3	2.022185	6.558874	0.31	0.759	-11.20504	15.24942
reg4	2.544457	7.271764	0.35	0.728	-12.12045	17.20937
pcturban	.0882531	.1900277	0.46	0.645	-.2949744	.4714805
_cons	28.19295	19.80007	1.42	0.162	-11.73771	68.1236

```

. testparm faminc reg2-reg4

( 1)  faminc = 0
( 2)  reg2 = 0
( 3)  reg3 = 0
( 4)  reg4 = 0

      F( 4,      43) =    12.16
      Prob > F =      0.0000

```

We can run some tests to determine if the exclusion restriction hold for our instruments if we have multiple instruments and assume that at least one of them is legitimate. The first test we will conduct is the LM test (or Sargan test). First estimate the IV.

```

. ivregress 2sls rent pcturban (hsngval = faminc reg2-reg4)

Instrumental variables (2SLS) regression               Number of obs   =          50
                                                       Wald chi2(2)    =          90.76
                                                       Prob > chi2     =          0.0000
                                                       R-squared       =          0.5989
                                                       Root MSE       =          22.166

```

rent	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
hsngval	.0022398	.0003284	6.82	0.000	.0015961 .0028836
pcturban	.081516	.2987652	0.27	0.785	-.504053 .667085
_cons	120.7065	15.22839	7.93	0.000	90.85942 150.5536

```

Instrumented:  hsngval
Instruments:   pcturban faminc reg2 reg3 reg4

```

Predict the residual from the IV.

```

. predict IVRES, res

```

Regress the residual on the controls and instruments.

```

. reg IVRES pcturban faminc reg2-reg4

      Source |           SS           df           MS       Number of obs   =          50
-----+-----+-----+-----+----- F(5, 44)      =          2.57
      Model |  5545.79184            5    1109.15837   Prob > F          =          0.0402
      Residual | 19019.9253           44    432.271031   R-squared         =          0.2258
-----+-----+-----+-----+----- Adj R-squared    =          0.1378
      Total | 24565.7172           49    501.341167   Root MSE         =          20.791

```

IVRES	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
pcturban	-.2827938	.2584246	-1.09	0.280	-.8036144 .2380269
faminc	.0044478	.0015321	2.90	0.006	.00136 .0075356
reg2	-6.768287	9.261742	-0.73	0.469	-25.4341 11.89753
reg3	4.847343	9.1507	0.53	0.599	-13.59468 23.28937
reg4	-18.76881	9.09554	-2.06	0.045	-37.09967 -.4379549
_cons	-62.8457	26.95197	-2.33	0.024	-117.1638 -8.52757

We need to calculate N*R2. Typing this command shows you all the results that are being held in temporary memory.

```

. ereturn list

```

```

scalars:
      e(N) = 50
      e(df_m) = 5
      e(df_r) = 44
      e(F) = 2.565886421248727
      e(r2) = .2257533047709994

```

```

e(rmse) = 20.79112865231076
e(mss) = 5545.791839052552
e(rss) = 19019.92534802524
e(r2_a) = .1377707257677039
e(ll) = -219.4774118038267
e(ll_0) = -225.8740300160578
e(rank) = 6

```

macros:

```

e(cmdline) : "regress IVRES pcturban faminc reg2-reg4"
e(title) : "Linear regression"
e(marginsok) : "XB default"
e(vce) : "ols"
e(depvar) : "IVRES"
e(cmd) : "regress"
e(properties) : "b V"
e(predict) : "regres_p"
e(model) : "ols"
e(estat_cmd) : "regress_estat"

```

matrices:

```

e(b) : 1 x 6
e(V) : 6 x 6

```

functions:

```

e(sample)

```

Here is our N*R2

```

. scalar nr2 = e(N)*e(r2)

. di nr2
11.287665

```

And the test statistic is distributed Chi-2 with degrees of freedom = # of overidentified instruments (in this case we have 4 instruments or 3 extra). The null is that there is no relationship between the residuals and the instruments, which in this case we reject. This means at least one of those instruments is bad.

```

. di 1 - chi2(3, nr2)
.01026785

```

A simpler way to do this is to estimate the IV.

```

. ivregress 2sls rent pcturban (hsngval = faminc reg2-reg4)

```

```

Instrumental variables (2SLS) regression      Number of obs   =          50
                                              Wald chi2(2)     =          90.76
                                              Prob > chi2      =          0.0000
                                              R-squared       =          0.5989
                                              Root MSE       =          22.166

```

rent	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
hsngval	.0022398	.0003284	6.82	0.000	.0015961 .0028836
pcturban	.081516	.2987652	0.27	0.785	-.504053 .667085
_cons	120.7065	15.22839	7.93	0.000	90.85942 150.5536

```

Instrumented:  hsngval
Instruments:   pcturban faminc reg2 reg3 reg4

```

Then use the estat overid command. The Sargan score is the LM test we just calculated.

```

. estat overid

```

Tests of overidentifying restrictions:

```

Sargan (score) chi2(3) = 11.2877 (p = 0.0103)
Basman chi2(3)       = 12.8294 (p = 0.0050)

```

We know that at least one of the instruments is problematic. To do this next test (another Hausman) we need to first make an assumption about which instrument we believe is legitimate and then test if the others are problematic. In this example I assumed that the family income instrument is legitimate. So, estimate the IV just using family income as an instrument.

```
. ivregress 2sls rent pcturban (hsngval = faminc)
```

Instrumental variables (2SLS) regression

Number of obs	=	50
Wald chi2(2)	=	51.44
Prob > chi2	=	0.0000
R-squared	=	0.2887
Root MSE	=	29.517

rent	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
hsngval	.0031938	.0006206	5.15	0.000	.0019775 .0044101
pcturban	-.5064118	.4815558	-1.05	0.293	-1.450244 .4374202
_cons	113.8143	20.52667	5.54	0.000	73.5828 154.0459

Instrumented: hsngval
Instruments: pcturban faminc

```
. est store my_IVe
```

And then compare the coefficients from the IV using only family income (this is our consistent estimator) to IV using all the instruments (this is our more efficient estimator that might be biased). The null is that there is no difference in the coefficients and we can use the full IV. The alternative is that the coefficients are different and we should only use family income because the extra instruments are problematic (correlated with the error). In this case we reject the null and conclude that we should not use extra instruments (census region). Just to reiterate, we have to assume that family income is a legitimate instrument to conduct this test.

```
. hausman my_IVe my_IV, sigmamore
```

Note: the rank of the differenced variance matrix (1) does not equal the number of coefficients being tested (2); be sure this is what you expect, or there may be problems computing the test. Examine the output of your estimators for anything unexpected and possibly consider scaling your variables so that the coefficients are on a similar scale.

---- Coefficients ----				
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	my_IVe	my_IV	Difference	S.E.
hsngval	.0031938	.0022398	.000954	.0003306
pcturban	-.5064118	.081516	-.5879278	.2037461

b = consistent under Ho and Ha; obtained from ivregress
B = inconsistent under Ha, efficient under Ho; obtained from ivregress

Test: Ho: difference in coefficients not systematic

```
chi2(1) = (b-B)'[(V_b-V_B)^(-1)](b-B)
        = 8.33
Prob>chi2 = 0.0039
```

```
. log close
```