

Webinar Series in Applied Quantitative Analysis

Date	Topic	
	<u>Session One</u>	
February 29	Potential Outcomes and Omitted Variable Bias I (Theory)	✓
March 7	Potential Outcomes and Omitted Variable Bias II (Application)	✓
	<u>Session Two</u>	
March 21	Difference-in-differences I (Theory)	✓
March 28	Difference-in-differences II (Application)	
	<u>Session Three</u>	
April 25	Power analysis, clustering and sample size calculations I (Theory)	
May 2	Power analysis, clustering and sample size calculations II (Application)	
	<u>Session Four</u>	
May 23	Fixed-effects I (Theory)	
May 30	Fixed-effects II (Application)	
	<u>Session Five</u>	
June 20	Instrumental variables I (Theory)	
June 27	Instrumental Variables II (Application)	
	<u>Session Six</u>	
July 25	Lagged dependent variables and the Arellano-Bond Estimator I (Theory)	
August 1	Lagged dependent variables and the Arellano-Bond Estimator II (Application)	



Advancing Evidence.
Improving Lives.

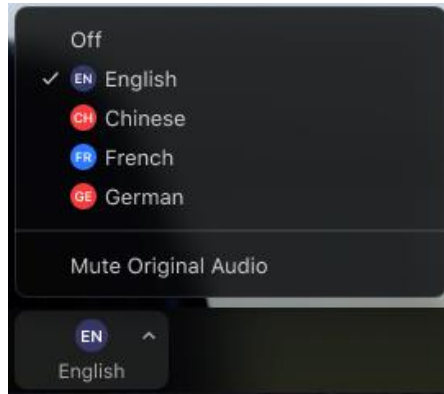
<https://www.airp3-africa.org/applied-quantitative-analysis-webinar-series>



Listening to language interpretation

Windows | macOS

1. In your meeting/webinar controls, click Interpretation .
2. Click the language that you would like to hear. (We will have French) No need to choose English, that is the language in the main Zoom room



3. (Optional) To hear the interpreted language only, click Mute Original Audio.

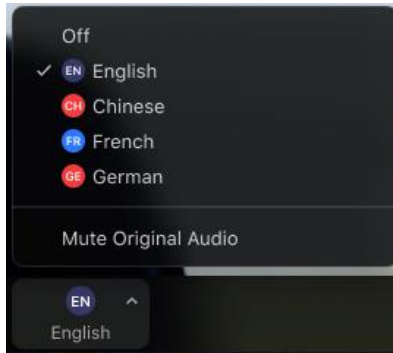
Notes:

- You must join the meeting audio through your computer audio/VoIP. You cannot listen to language interpretation if you use the dial-in or call me phone audio features.
- As a participant joining a language channel, you can broadcast back into the main audio channel if you unmute your audio and speak.

Écoute de l'interprétation d'une langue

Windows | macOS

1. Dans les contrôles de votre réunion/webinaire, cliquez sur Interprétation .
2. Cliquez sur la langue que vous souhaitez entendre. (Nous aurons le français) Pas besoin de choisir l'anglais, c'est la langue de la salle Zoom principale



3. (Facultatif) Pour entendre uniquement la langue interprétée, cliquez sur Couper le son original.

Remarques:

- Vous devez rejoindre l'audio de la réunion via l'audio/VoIP de votre ordinateur. Vous ne pouvez pas écouter l'interprétation linguistique si vous utilisez les fonctions audio de connexion ou d'appel téléphonique.
- En tant que participant rejoignant une chaîne linguistique, vous pouvez retransmettre sur le canal audio principal canal si vous réactivez votre audio et parlez.

Difference-in-Differences II

Ashu Handa

Institute Fellow – AIR

Kenan Eminent Professor of Public Policy – UNC-CH

March 28, 2024

[Some slides are courtesy of Prof. Gustavo Angeles]

Topics for today

- Review the 2x2 DiD model and answer lingering questions
- Extend the model to additional time periods
- Brief intro to DiD with staggered treatment (advanced topic)
- STATA demonstrations (data and scripts are on the website)

$$Y = B_0 + B_1 * (\text{Treat}) + B_2 * (\text{Post}) + B_3 * (\text{Treat} * \text{Post}) + e_i$$

Treat=1 if treated group, 0 if control group

Post=1 if post-intervention period (after), 0 if pre-intervention

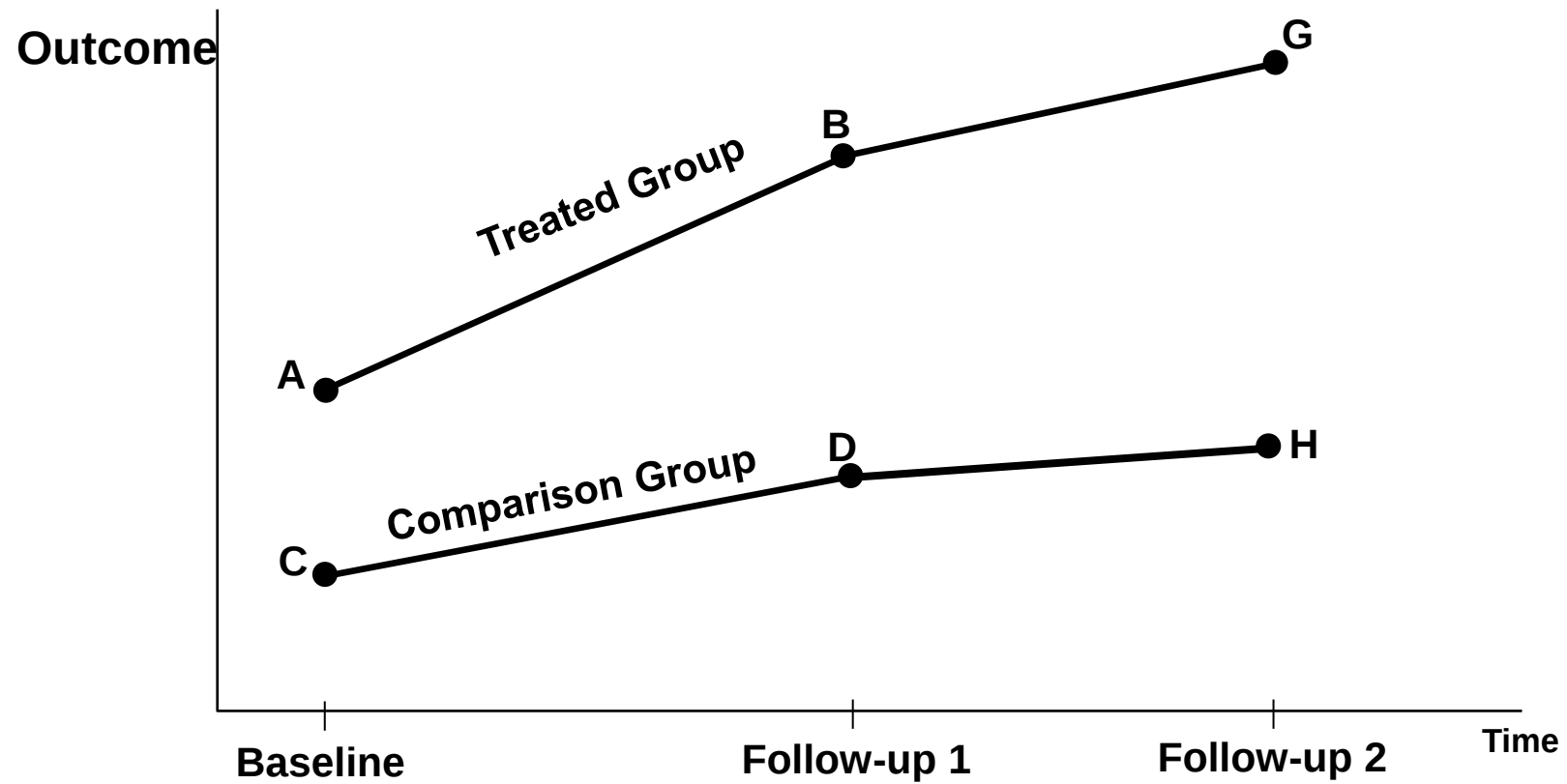
	Before (Post=0)	After (Post=1)	Difference
Comparison (Treat=0)	$Y = B_0$	$Y = B_0 + B_2$	B_2
Treated (Treat=1)	$Y = B_0 + B_1$	$Y = B_0 + B_1 + B_2 + B_3$	$B_2 + B_3$
Difference	B_1	$B_1 + B_3$	B_3

← DID Impact

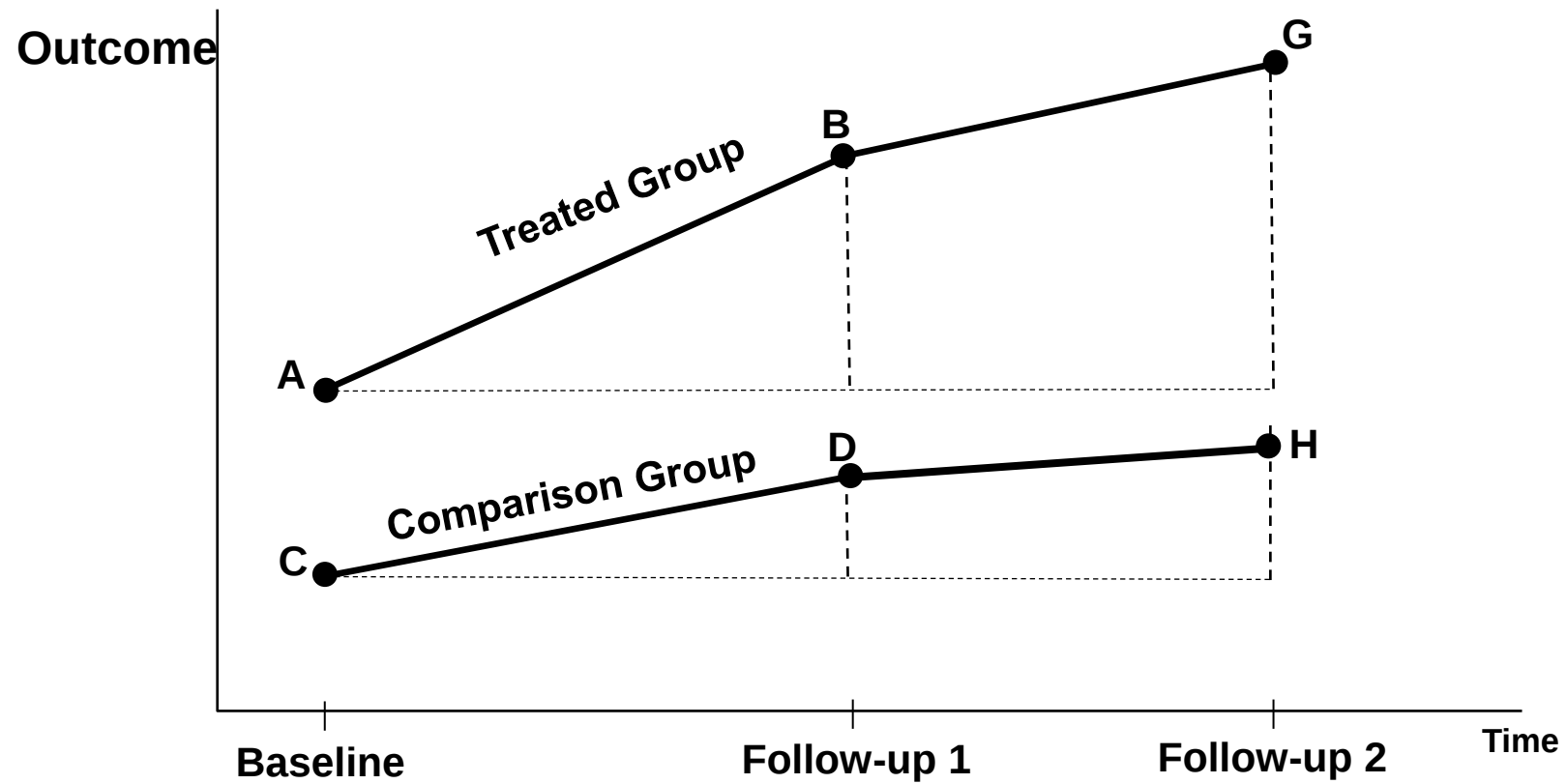
Coefficient interpretation

1. B_2 is the time trend
2. B_1 gives the difference between the two groups before the intervention
3. B_0 is the mean of Y in the comparison group before the intervention
4. B_3 is the DiD estimate

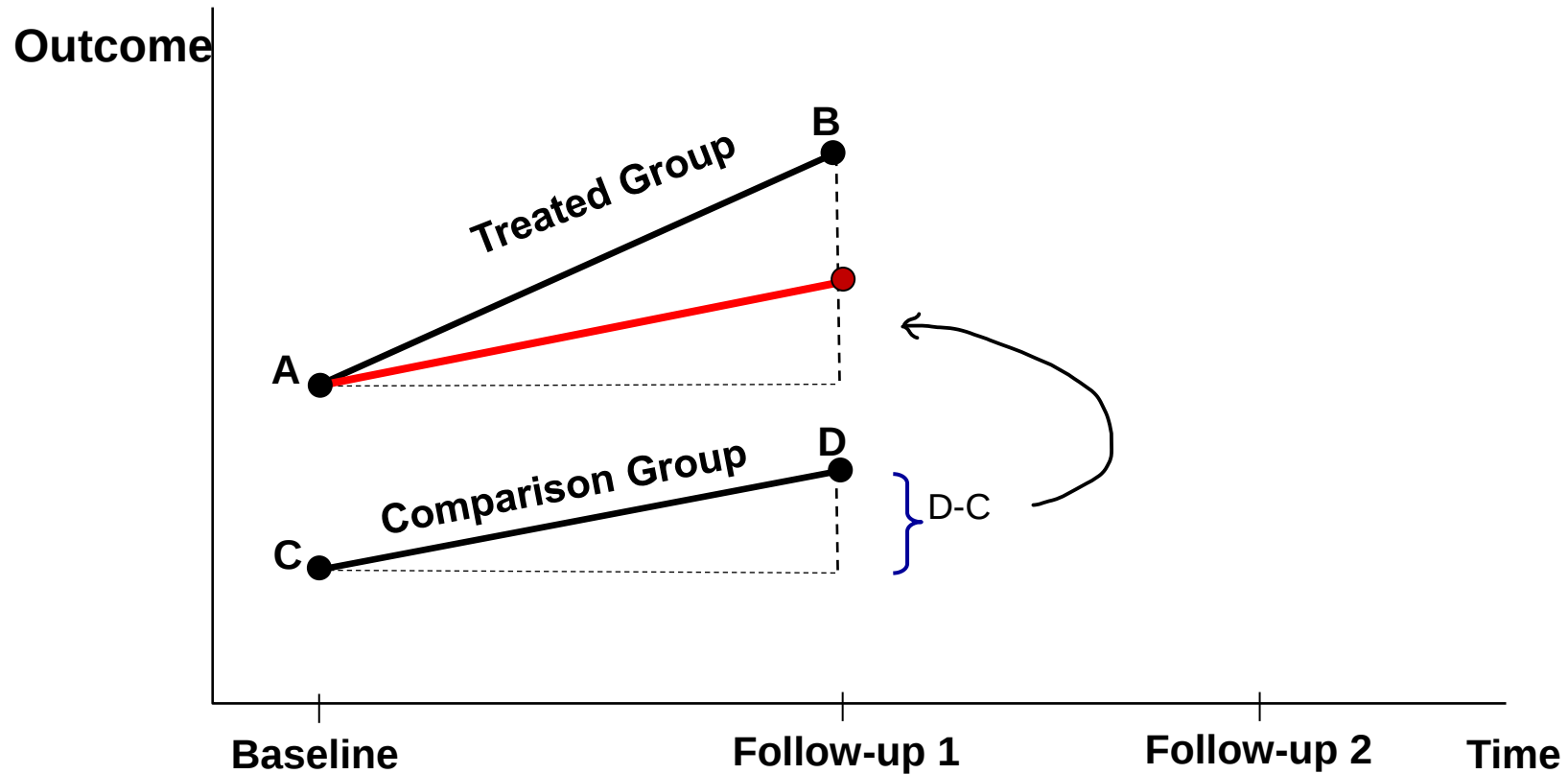
DID Extensions: Multiple times – Case: 3 points in time



DiD Extensions: Multiple times – Case: 3 points in time



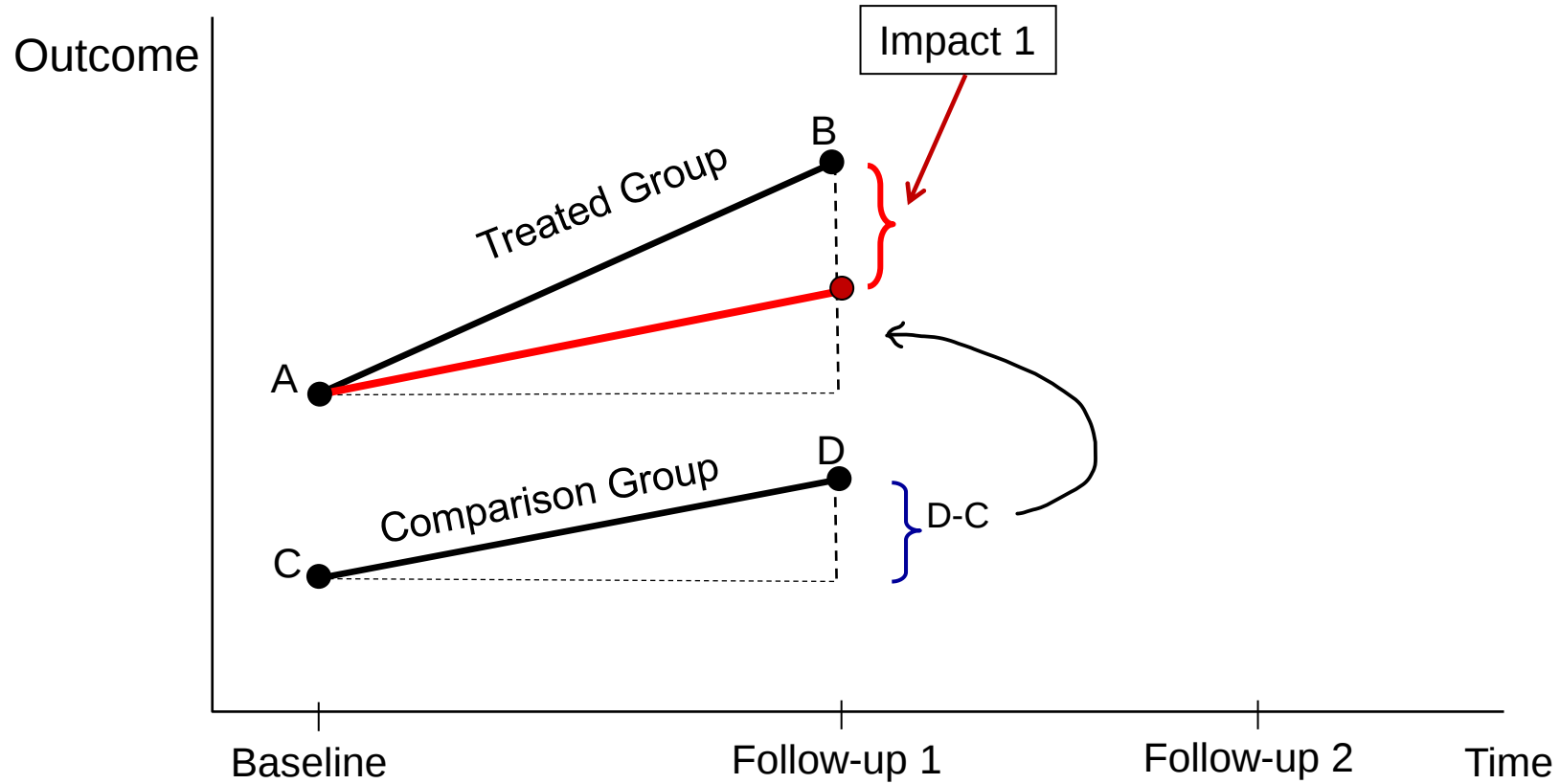
DiD Extensions - 3 points in time



Key assumption 1: Parallel trends assumption holds for each time period.

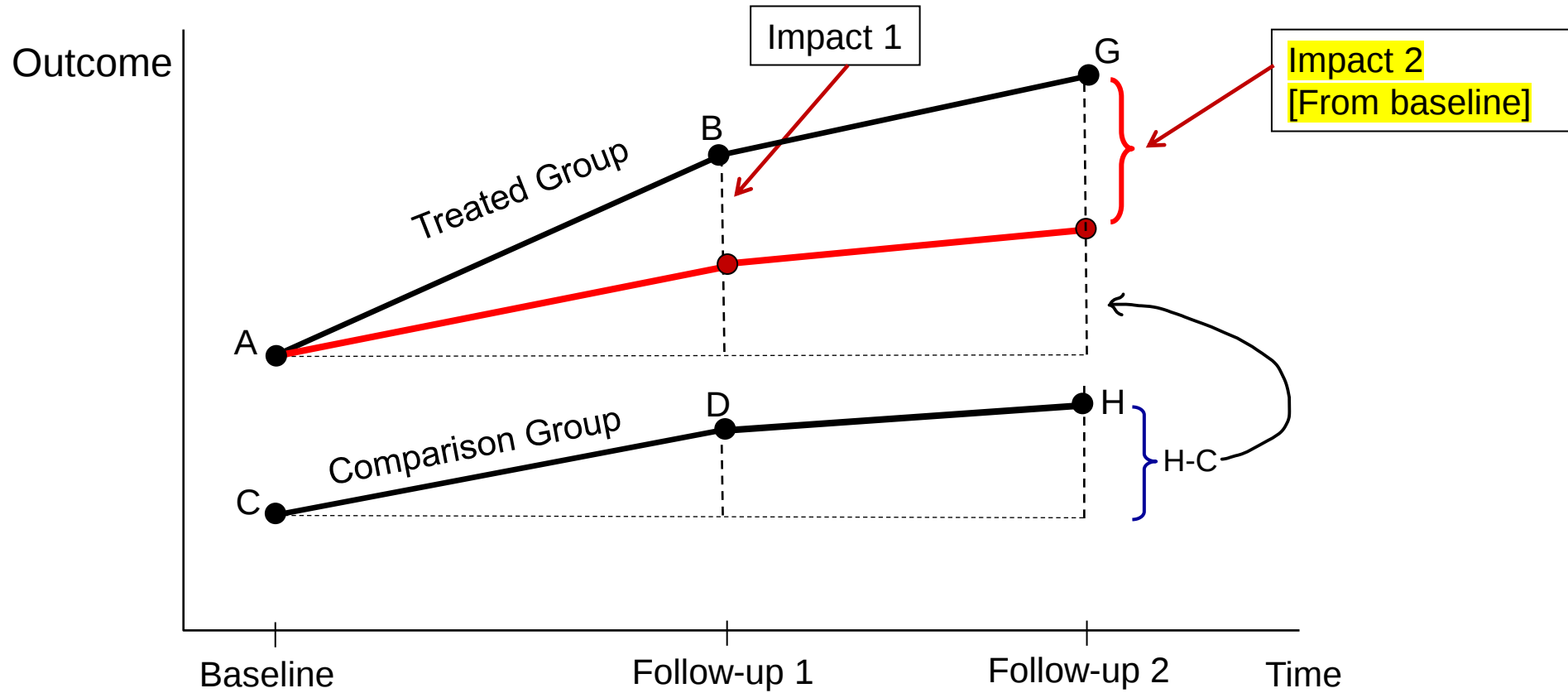
Key assumption 2: No anticipation effects

DiD Extensions - 3 points in time



Key assumption 1: Parallel trends assumption holds for each time period.

DiD Extensions - 3 points in time



Key assumption 1: Parallel trends assumption holds for each time period.

DID Extensions - 3 points in time: The empirical model

The model:

$$Y_{ijt} = \beta_0 + \beta_1 \text{Treat}_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 \text{Treat}_j * T1_t + \beta_5 \text{Treat}_j * T2_t + \beta_6 X_{ijt} + \varepsilon_{ijt}$$

Where,

Y_{ijt} : Outcome of interest

$\text{Treat}_j = \begin{cases} 1, & \text{if community } j \text{ is in the treated area} \\ 0, & \text{if community } j \text{ is in the comparison area} \end{cases}$

$T1_t = \begin{cases} 1, & \text{if observation is from Follow – up 1 survey} \\ 0, & \text{if observation is from other time} \end{cases}$

$T2_t = \begin{cases} 1, & \text{if observation is from Follow – up 2 survey} \\ 0, & \text{if observation is from other time} \end{cases}$

X_{ijt} : Individual characteristics

ε_{ijt} : Random error

DID Extensions - 3 points in time: The empirical model

The model:

$$Y_{ijt} = \beta_0 + \beta_1 \text{Treat}_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 \text{Treat}_j * T1_t + \beta_5 \text{Treat}_j * T2_t + \beta_6 X_{ijt} + \varepsilon_{ijt}$$

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DID Extensions - 3 points in time: The empirical model

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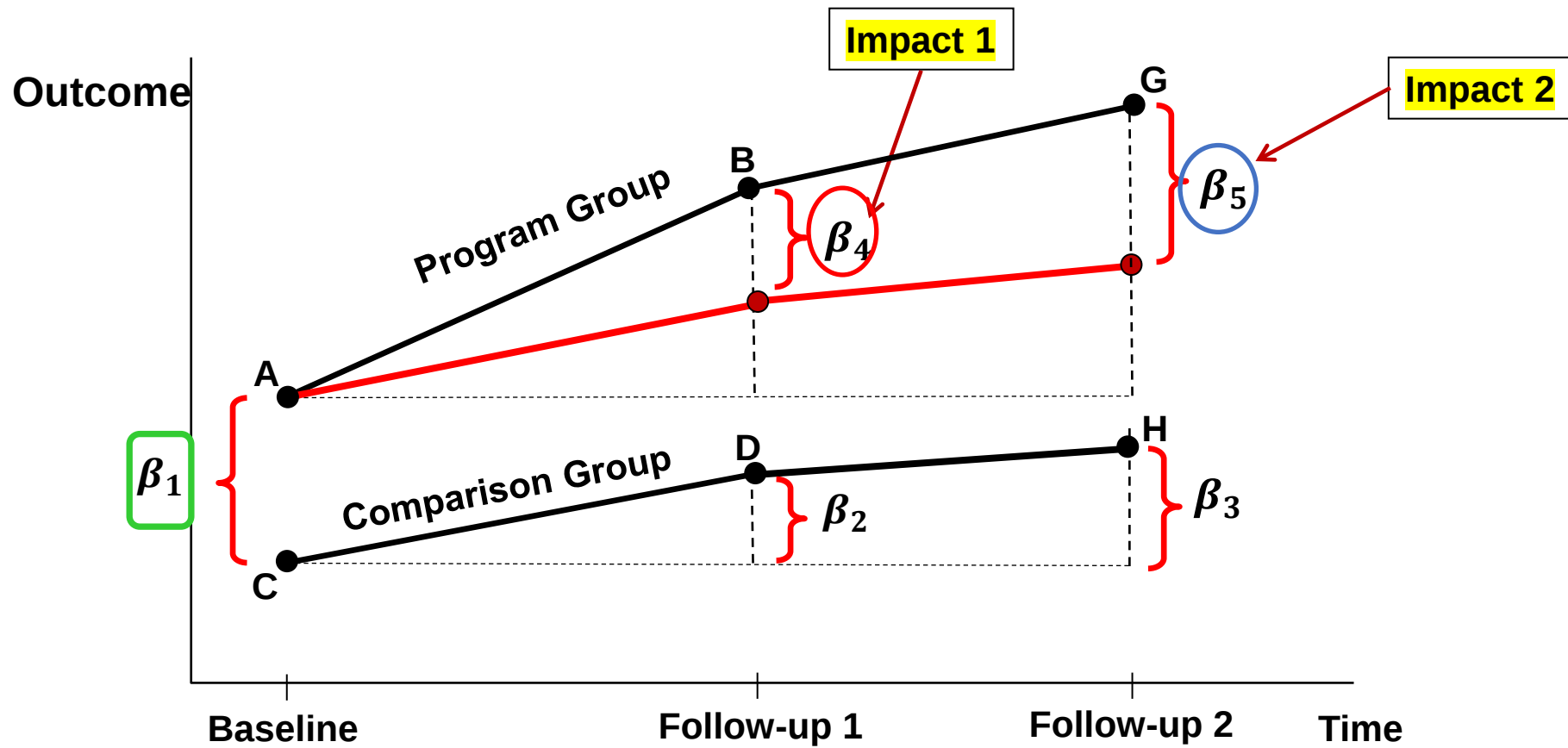
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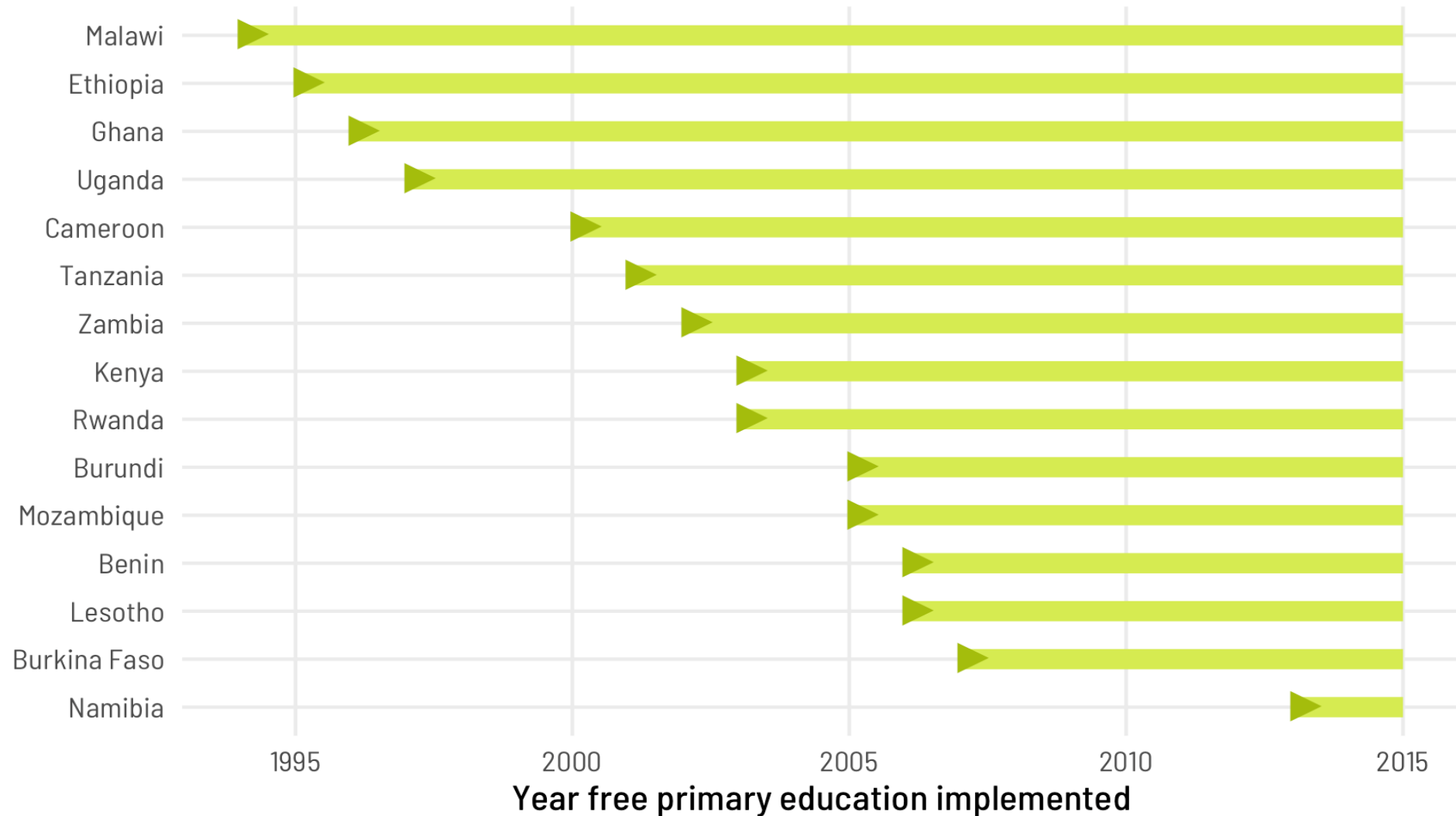


Assumption 1: Parallel trends assumption holds for each time period.

New developments in DiD that address deviations from the standard 2x2 model (brief overview)

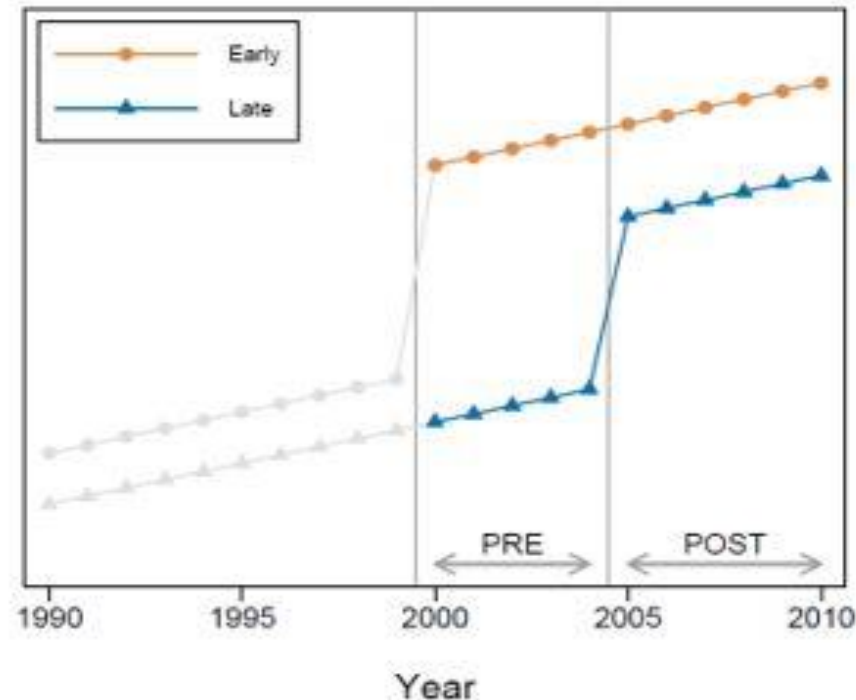
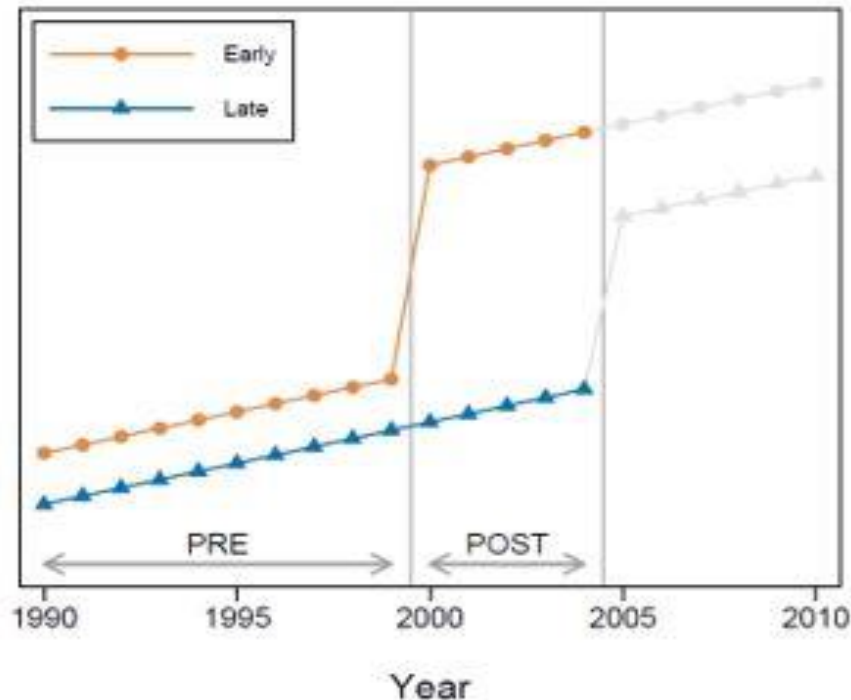
1. Multiple post-intervention periods with staggered treatment;
2. Multiple post-intervention periods with heterogeneous treatment effects over time;
3. Treatment heterogeneity across units (where units refer to the level of treatment—clusters, countries, districts);
4. A combination of #2 and #3.

Effect of free primary on schooling: staggered treatment



Intuition: Effect of free primary education on schooling

- **Two types of countries:** one group that made primary education free in 2000 (early), and another group that made it free in 2005 (late)
- **Data:** Primary school completion rates every year from 1990 to 2010.
- **Two distinct DiD comparisons are possible:**
 - First, we could focus on the period from 1990 to 2004 (**the left panel in the figure below**) - late adopter countries are “never treated” – so they can be used as a comparison group to estimate the impact of free primary schooling on test scores in the early adopter countries
 - Second, we can also construct a second DD estimate of the treatment effect of free primary school by focusing on the years 2001 to 2010 (**the right panel in the figure below**). Treatment status of the early adopters never changes – they remain treated throughout – so they can be used as a comparison group to estimate the impact of free primary schooling on test scores in the late adopting countries



Key references if your study deviates from 2x2

[Red papers have ready-made code you can use!]

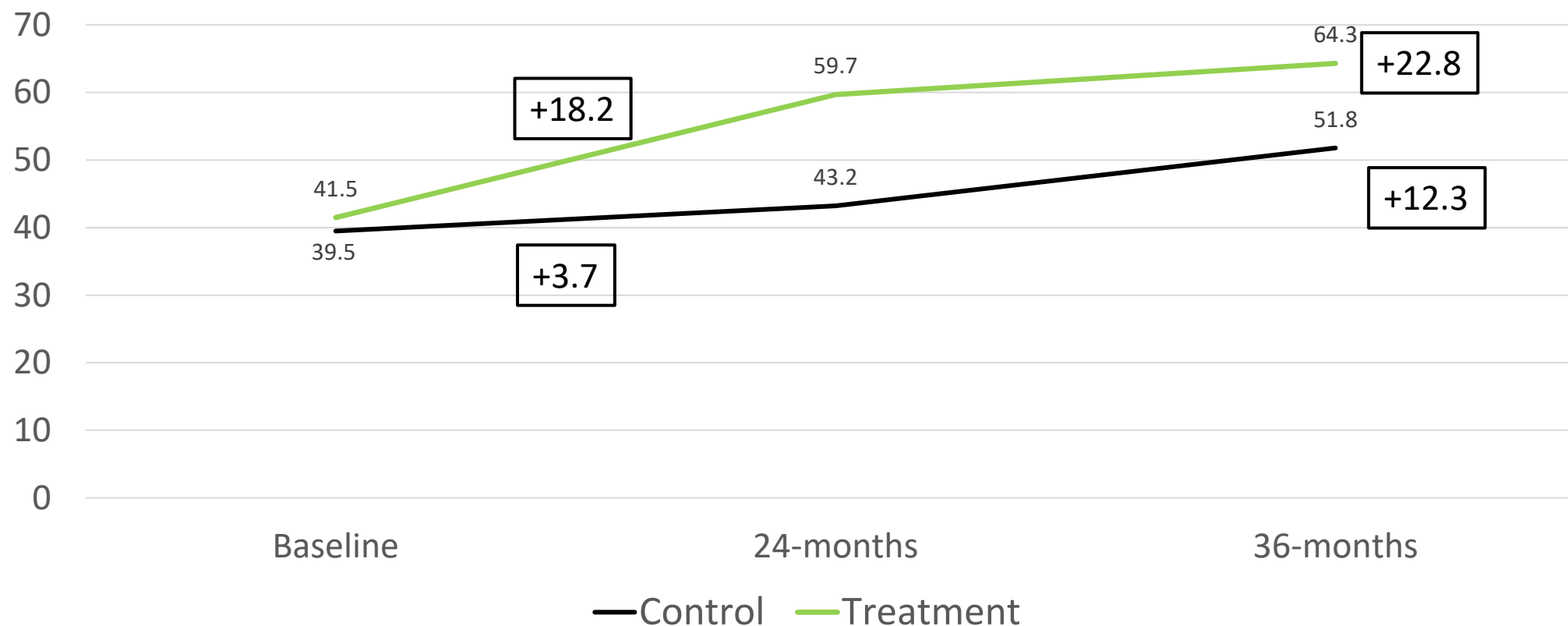
- Baker, Andrew, David F. Larcker, and Charles C. Y. Wang. 2021. “How Much Should We Trust Staggered Difference-in-Differences Estimates?” Working paper 246. Rock Center for Corporate Governance at Stanford University. <https://doi.org/10.2139/ssrn.3794018>.
- Callaway, Brantly, and Pedro H. C. Sant’Anna. 2020. “Difference-in-Differences with Multiple Time Periods.” *Journal of Econometrics*, December. <https://doi.org/10.1016/j.jeconom.2020.12.001>.
- de Chaisemartin, Clément, and Xavier d’Haultfoeuille. 2020. “Two-Way Fixed Effects Estimators with Heterogeneous Treatment Effects.” *American Economic Review* 110 (9): 2964–96. <https://doi.org/10.1257/aer.20181169>.
- Goodman-Bacon, Andrew. 2021. “Difference-in-Differences with Variation in Treatment Timing.” *Journal of Econometrics*, June. <https://doi.org/10.1016/j.jeconom.2021.03.014>.
- Jakiela, Pamela. 2021. “Simple Diagnostics for Two-Way Fixed Effects,” March. <https://arxiv.org/abs/2103.13229>.
- Sun, Liyang, and Sarah Abraham. 2020. “Estimating Dynamic Treatment Effects in Event Studies with Heterogeneous Treatment Effects.” *Journal of Econometrics*, December. <https://doi.org/10.1016/j.jeconom.2020.09.006>.
- <https://www.andrewheiss.com/blog/2021/08/25/twfe-diagnostics/#a-different-way-of-thinking-about-ols-coefficients>

STATA demonstrations

- Zambia Child Grant
 - 2x2 DiD
 - Three periods
- Rural-urban differences in male youth outcomes over time in Ghana
 - Repeated cross-sections using DHS
 - Data is 'stacked'
 - 'Treatment' is 'living in rural region'
 - **Hundreds of PhD dissertations available using this methodology**
- Measuring effect of climate shocks (time permitting)

Zambia Child Grant Program Impact Evaluation

Per Capita Consumption (Zambian Kwacha)



DiD estimates

Baseline-24-months: $(18.2 - 3.7) = +14.5$

Baseline-36-months: $(22.8 - 12.3) = +10.5$ (from baseline)

$$Y_{ijt} = \beta_0 + \beta_1 Treat_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 Treat_j * T1_t + \beta_5 Treat_j * T2_t + \varepsilon_{ijt}$$

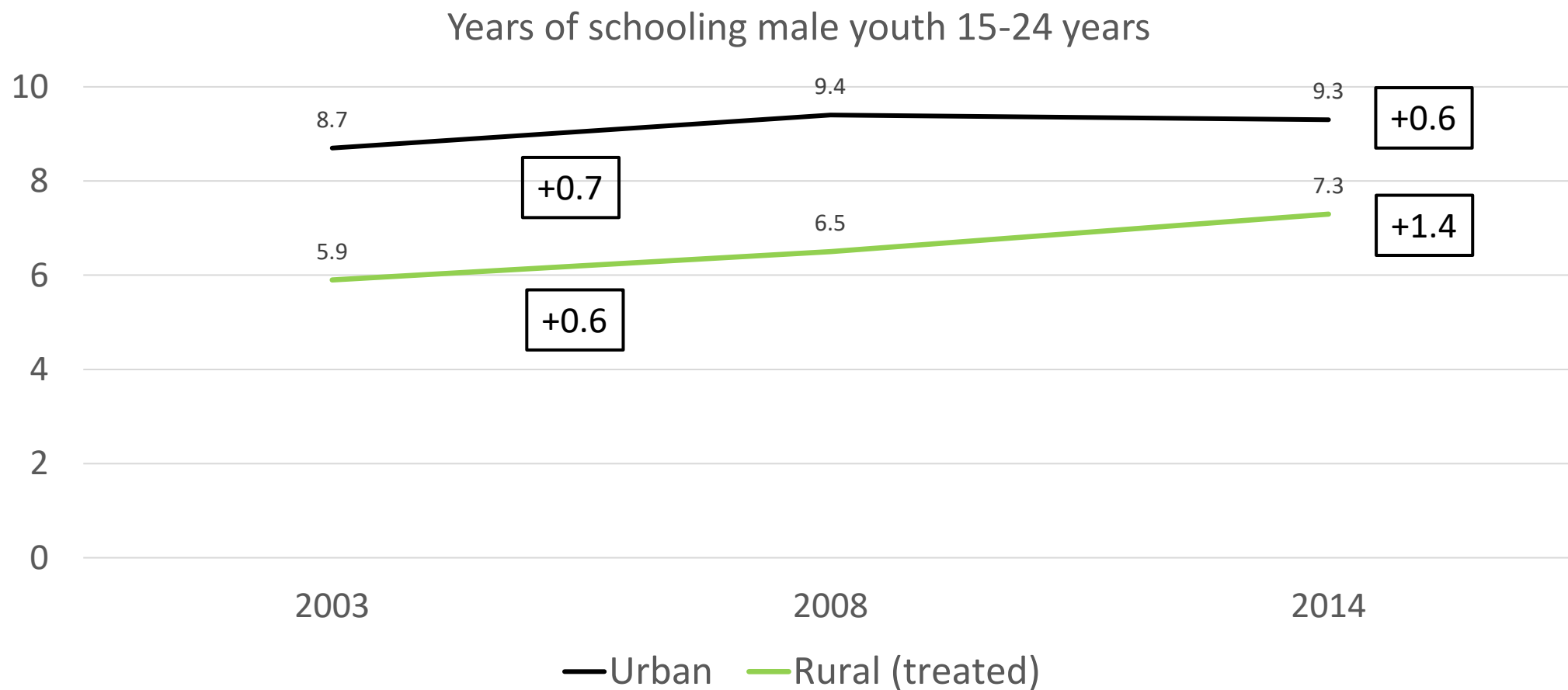
39.5	2.0	3.8	12.4	14.4	10.4
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```
. reg exp_per_capita treat period24 period36 DD24 DD36
```

Source	SS	df	MS	Number of obs	=	7,273
				F(5, 7267)	=	97.26
Model	644287.314	5	128857.463	Prob > F	=	0.0000
Residual	9627429.76	7,267	1324.81488	R-squared	=	0.0627
				Adj R-squared	=	0.0621
Total	10271717.1	7,272	1412.50235	Root MSE	=	36.398

exp_per_ca~a	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
treat	2.019155	1.451573	1.39	0.164	-.8263499	4.86466
period24	3.789231	1.486378	2.55	0.011	.8754988	6.702963
period36	12.35714	1.45655	8.48	0.000	9.501876	15.2124
DD24	14.40795	2.100744	6.86	0.000	10.28988	18.52602
DD36	10.44709	2.064304	5.06	0.000	6.400455	14.49373
_cons	39.47611	1.025805	38.48	0.000	37.46524	41.48699

Rural-urban differences in youth over time in Ghana



DiD estimates

2003-2008: $(0.6 - +0.7) = -0.1$

2003-2014: $(1.4 - +0.6) = 0.8$

$$Y_{ijt} = \beta_0 + \beta_1 Treat_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 Treat_j * T1_t + \beta_5 Treat_j * T2_t + \varepsilon_{ijt}$$

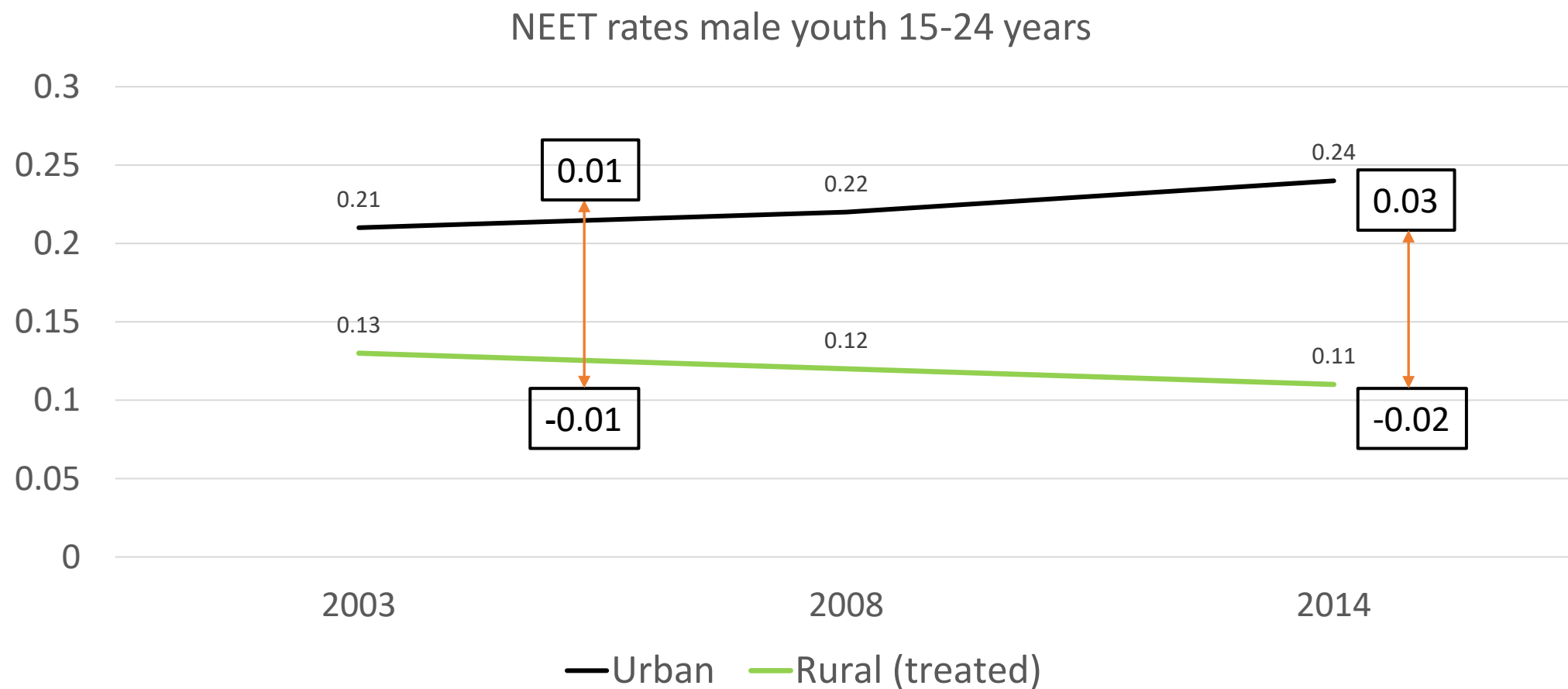
8.6	-2.8	0.69	0.58	-0.04	0.80
-----	------	------	------	-------	------

```
. reg yearsed year2008 year2014 rural DD2008 DD2014
```

Source	SS	df	MS	Number of obs	=	4,940
Model	9241.59087	5	1848.31817	F(5, 4934)	=	169.03
Residual	53952.6073	4,934	10.9348616	Prob > F	=	0.0000
				R-squared	=	0.1462
				Adj R-squared	=	0.1454
Total	63194.1982	4,939	12.7949379	Root MSE	=	3.3068

yearsed	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
year2008	.6932637	.1752934	3.95	0.000	.3496106	1.036917
year2014	.5824653	.1767417	3.30	0.001	.235973	.9289577
rural	-2.804055	.1592733	-17.61	0.000	-3.116301	-2.491808
DD2008	-.0387703	.2293317	-0.17	0.866	-.4883624	.4108219
DD2014	.8023604	.2338181	3.43	0.001	.343973	1.260748
_cons	8.674931	.1227265	70.69	0.000	8.434333	8.91553

Rural-urban differences in youth over time in Ghana



DiD estimates

2003-2008: $(-0.01 - + 0.01) = -0.02$

2003-2014: $(-0.02 - + 0.03) = -0.05$

$$Y_{ijt} = \beta_0 + \beta_1 Treat_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 Treat_j * T1_t + \beta_5 Treat_j * T2_t + \varepsilon_{ijt}$$

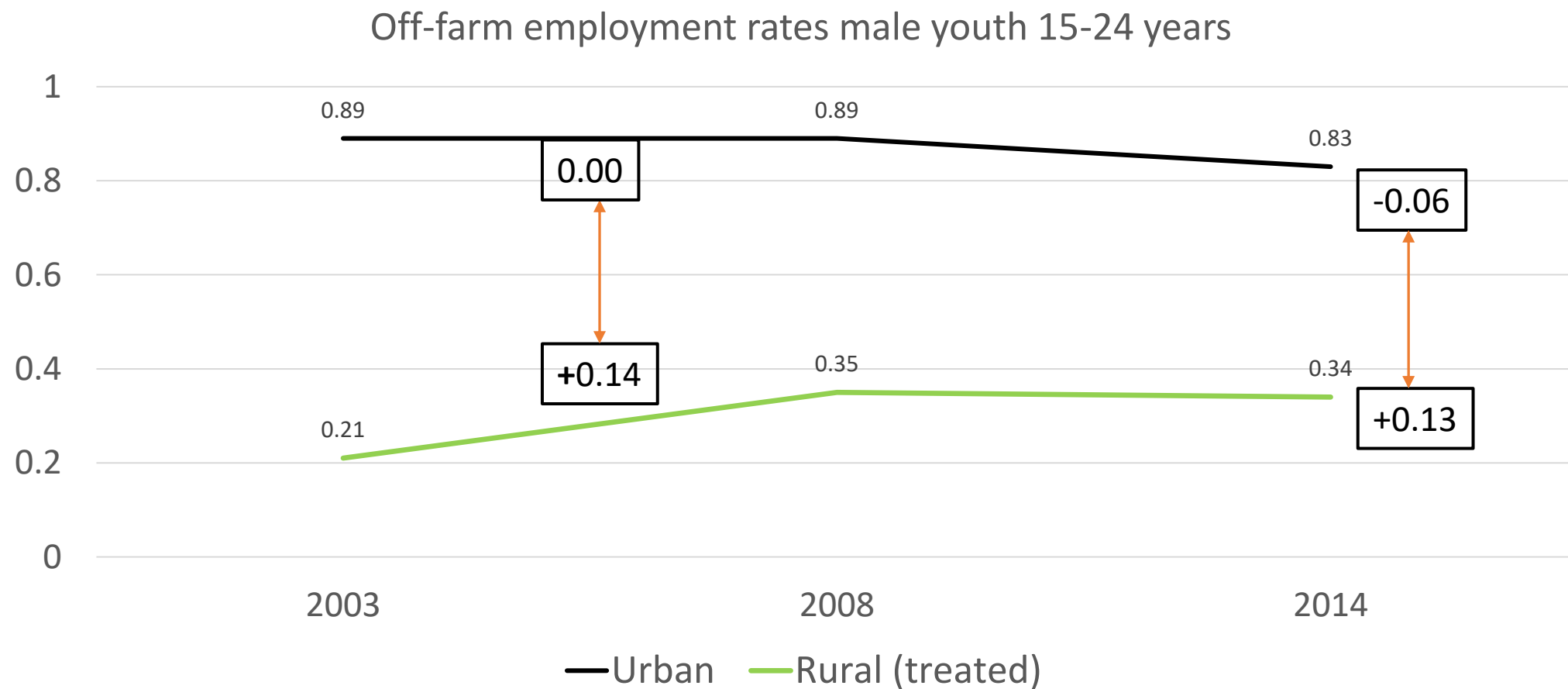
0.21	-0.08	0.02	0.03	-0.03	-0.05
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```
. reg neet year2008 year2014 rural DD2008 DD2014
```

Source	SS	df	MS	Number of obs	=	4,944
				F(5, 4938)	=	20.46
Model	13.7062056	5	2.74124111	Prob > F	=	0.0000
Residual	661.568673	4,938	.133975025	R-squared	=	0.0203
				Adj R-squared	=	0.0193
Total	675.274879	4,943	.136612357	Root MSE	=	.36603

neet	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
year2008	.0162456	.019389	0.84	0.402	-.0217654	.0542565
year2014	.0345127	.0195634	1.76	0.078	-.0038403	.0728656
rural	-.0784306	.0176298	-4.45	0.000	-.1129929	-.0438683
DD2008	-.0273379	.0253678	-1.08	0.281	-.0770702	.0223943
DD2014	-.05345	.0258811	-2.07	0.039	-.1041885	-.0027114
_cons	.2066116	.0135845	15.21	0.000	.1799799	.2332432

Rural-urban differences in youth over time in Ghana



DiD estimates

2003-2008: $(+0.14 - 0.0) = +0.14$

2003-2014: $(+0.13 - -0.06) = +0.19$

$$Y_{ijt} = \beta_0 + \beta_1 Treat_j + \beta_2 T1_t + \beta_3 T2_t + \beta_4 Treat_j * T1_t + \beta_5 Treat_j * T2_t + \varepsilon_{ijt}$$

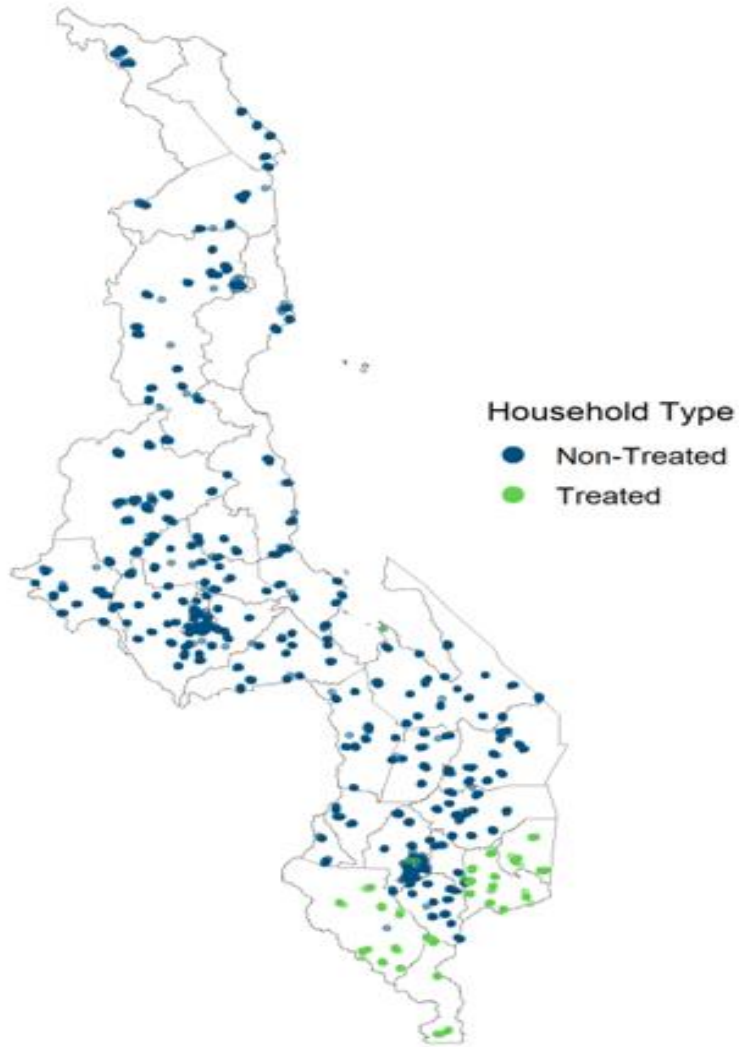
0.89	-0.68	0.01	-0.06	0.13	0.18
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```
. reg offfarm year2008 year2014 rural DD2008 DD2014
```

Source	SS	df	MS	Number of obs	=	2,539
				F(5, 2533)	=	222.83
Model	193.908114	5	38.7816227	Prob > F	=	0.0000
Residual	440.837062	2,533	.174037529	R-squared	=	0.3055
				Adj R-squared	=	0.3041
Total	634.745175	2,538	.250096602	Root MSE	=	.41718

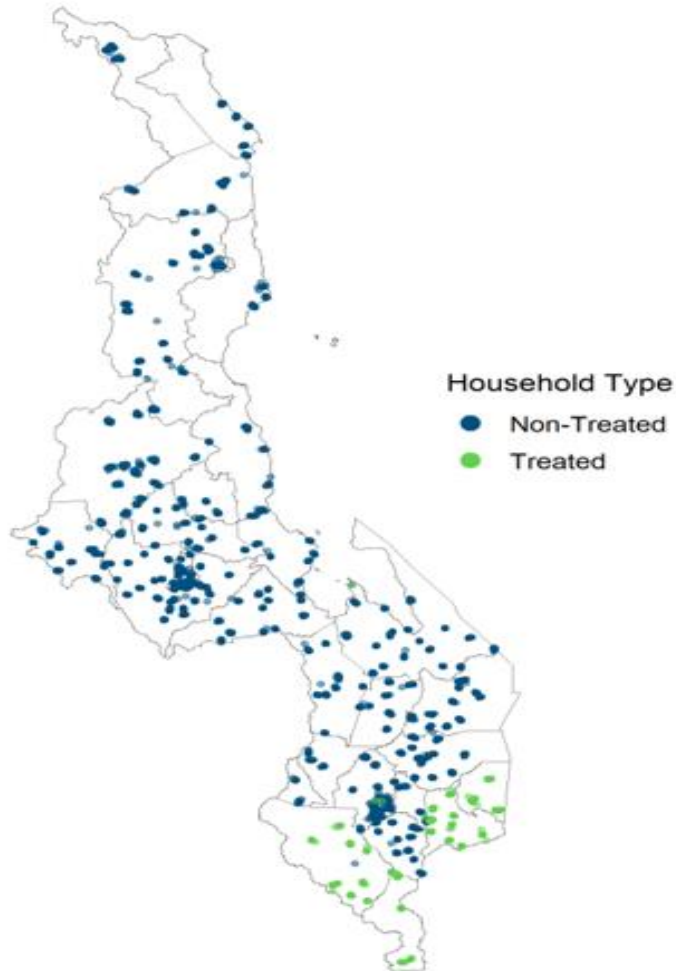
offfarm	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
year2008	.005668	.0352151	0.16	0.872	-.0633854	.0747214
year2014	-.0578935	.0351686	-1.65	0.100	-.1268556	.0110686
rural	-.6751369	.0320546	-21.06	0.000	-.7379927	-.6122811
DD2008	.133893	.0432617	3.09	0.002	.0490612	.2187249
DD2014	.1834484	.0434029	4.23	0.000	.0983396	.2685572
_cons	.8866397	.0265444	33.40	0.000	.8345887	.9386906

What was the 'impact' of Tropical Storms Ana and Gombe (Q1, 2022) on households in Malawi?



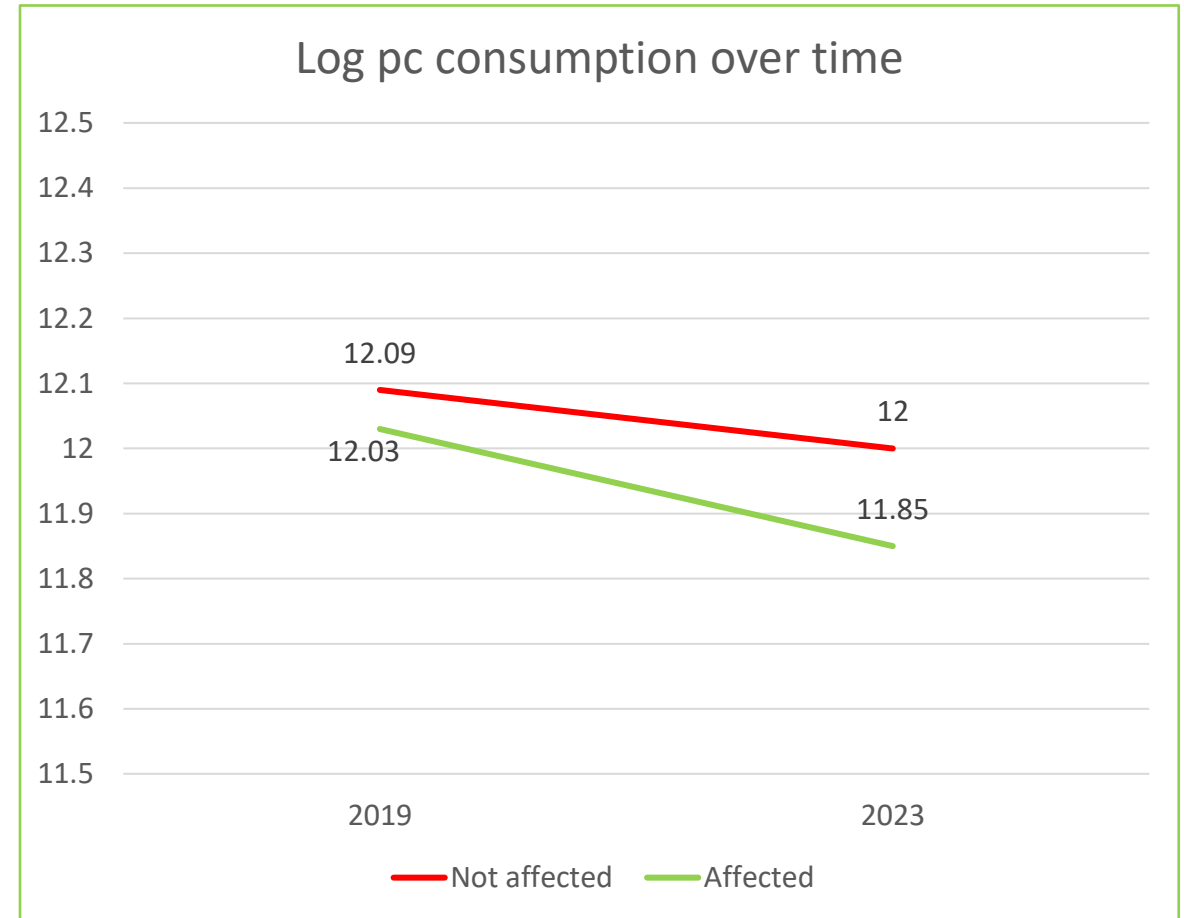
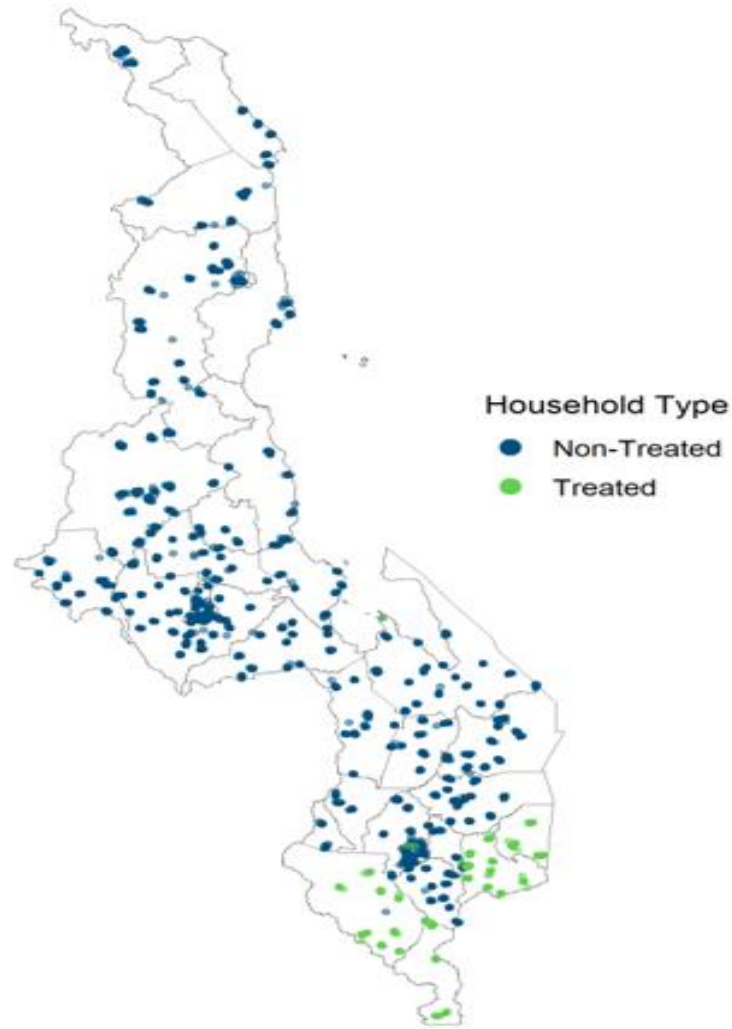
1. The 'treatment' is the storm locations (areas marked in green)
2. National household survey 2019 (IHS5)
3. Special national survey conducted in 2023

What was the ‘impact’ of Tropical Storms Ana and Gombe on households in Malawi?



1. Identified households in most intensely hit areas
2. Found ‘similar’ households in less intensely hit areas in the South
3. Had data from 2019 (before) and after (2023)—compared trends

DiD estimate of 'impact' of Tropical Storms Ana and Gombe in Malawi



Affected: $(12 - 12.09) = -0.09$

Not affected: $(11.85 - 12.03) = -0.18$

DiD: $(-0.18 - (-0.09)) = -0.09$ (9% reduction)

DiD estimate of 'impact' of Tropical Storms Ana and Gombé in Malawi

```
. reg logexp time T DD if bl_rural==1
```

Source	SS	df	MS	Number of obs	=	6,170
Model	23.1011513	3	7.70038376	F(3, 6166)	=	17.84
Residual	2661.79059	6,166	.431688386	Prob > F	=	0.0000
				R-squared	=	0.0086
				Adj R-squared	=	0.0081
Total	2684.89174	6,169	.435223171	Root MSE	=	.65703

logexp	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
time	-.0780356	.0179186	-4.35	0.000	-.1131623	-.0429088
T	-.0531358	.0353647	-1.50	0.133	-.1224629	.0161913
DD	-.1089284	.0500132	-2.18	0.029	-.2069718	-.0108851
_cons	12.08734	.0126704	953.98	0.000	12.0625	12.11218